A1. To choose a full-house hand, we must choose a rank for three cards $\binom{13}{1}$ possibilities), three cards in that rank $\binom{4}{3}$, one of the other ranks for two cards $\binom{12}{1}$, and two cards in that rank $\binom{4}{2}$. Therefore the probability is

$$\frac{\binom{13}{1}\binom{4}{3}\binom{12}{1}\binom{4}{2}}{\binom{52}{5}}.$$

A2. If we have four cards forming two pairs, then there are 48 remaining cards in the deck, and 4 of those cards will complete our hand to a full house. Therefore the probability is 4/48 = 1/12.

B1. [We discussed this problem recently in class, because it came up in the practice problems.] Let G_1 be the number of trials needed for the first success, G_2 the number of additional trials (after the first G_1 trials) needed for the second success, G_3 the number of additional trials (after the first $G_1 + G_2$ trials) needed for the third success, etc. Then $X = G_1 + \cdots + G_r$. Each G_k is a geometric random variable. The G_k are independent, because the geometric distribution is memoryless (in other words, because the underlying Bernoulli trials are independent).

B2. In class we showed that the MGF for each G_k is

$$m_{G_k}(t) = \frac{pe^t}{1 + e^t(p-1)}$$

Then, because the G_k are independent and X is their sum,

$$m_X(t) = \left(\frac{pe^t}{1 + e^t(p-1)}\right)^r.$$

C. Let *D* be the event that I have the disease and *T* the event that I test positive for it. We are told that $P(T|D) = P(T^c|D^c) = 0.95$. Also, P(D) = 0.01 is the probability that I have the disease before receiving any test results. We wish to find

$$P(D|T) = \frac{P(T|D)P(D)}{P(T|D)P(D) + P(T|D^{c})P(D^{c})}$$

= $\frac{0.95 \cdot 0.01}{0.95 \cdot 0.01 + 0.05 \cdot 0.99}$
= $\frac{95 \cdot 1}{95 \cdot 1 + 5 \cdot 99}$
= $95/590$
= $19/118.$

[By the way, that's about 16%.]

D. We begin by computing the CDF of Z:

$$F_Z(z) = P(Z \le z)$$

= $P(X + Y \le z)$
= $\iint_R f(x, y) \, dx \, dy,$

where R is the set of all points (x, y) such that $x + y \le z$. [A good solution is accompanied by a diagram showing R and thus explaining the following integration bounds. I omit the diagram in these typed solutions.] The integral equals

$$\int_{-\infty}^{\infty} \int_{-\infty}^{z-x} f(x,y) \, dy \, dx.$$

Therefore the density of Z is

$$f_Z(z) = \frac{d}{dz} F_Z(z)$$

= $\frac{d}{dz} \int_{-\infty}^{\infty} \int_{-\infty}^{z-x} f(x, y) \, dy \, dx$
= $\int_{-\infty}^{\infty} \frac{d}{dz} \int_{-\infty}^{z-x} f(x, y) \, dy \, dx$
= $\int_{-\infty}^{\infty} f(x, z - x) \, dx$

by the fundamental theorem of calculus.

E. Let $S_n = X_1 + \cdots + X_n$ as usual. We wish to find n such that

$$0.99 = P\left(\left|\frac{S_n}{n} - \mu\right| \le \epsilon\right)$$
$$= P\left(-\epsilon \le \frac{S_n}{n} - \mu \le \epsilon\right)$$
$$= P\left(-\frac{\epsilon}{\sigma/\sqrt{n}} \le Z \le \frac{\epsilon}{\sigma/\sqrt{n}}\right),$$

where

$$Z = \frac{\frac{S_n}{n} - \mu}{\sigma/\sqrt{n}}$$

is a standard normal random variable (approximately, for large n, by the central limit theorem). Let F be the CDF of Z. The probability above equals 0.99 when $F(\epsilon \sqrt{n}/\sigma) = 0.995$. Therefore the minimum usable n is

$$n = \left(\frac{\sigma}{\epsilon} F^{-1}(0.995)\right)^2.$$

[In practice, σ would be estimated from the data, and ϵ (and the 0.99) would be chosen according to the biologist's tastes and needs.]

F. Recall from class that the CDF of X is $F(x) = 1 - e^{-\lambda x}$. Setting u = F(x) and solving for x as a function of u gives

$$x = -\frac{1}{\lambda}\log(1-u).$$

So we choose u uniformly on [0, 1] and compute x from u using that equation.

G1. Well,

$$F_{X|X>s}(x|x>s) = P(X \le x|X>s)$$

$$= \frac{P(X \le x, X>s)}{P(X>s)}$$

$$= \frac{\int_s^x f(t) dt}{1 - F(s)}$$

$$\Rightarrow f_{X|X>s}(x|x>s) = \frac{d}{dx}F_{X|X>s}(x|x>s)$$

$$= \frac{\frac{d}{dx}\int_s^x f(t) dt}{1 - F(s)}$$

$$= \frac{f(x)}{1 - F(s)}$$

by the fundamental theorem of calculus. By the way, the domain of $f_{X|X>s}(x|x>s)$ is (s,∞) rather than $(-\infty,\infty)$. In other words, $f_{X|X>s}(x|x>s) = 0$ for $x \leq s$. This claim should make intuitive sense. You can also verify it by computing

$$\int_{s}^{\infty} f_{X|X>s}(x|x>s) \, dx = \frac{1}{1-F(s)} \int_{s}^{\infty} f(x) \, dx = \frac{1}{1-F(s)} (1-F(s)) = 1.$$

G2. [Because I made an error while writing this problem, it requires more computation than I intended. Grading was generous.] We are asked to compute

$$E[X|X>1] = \int_{1}^{10} x f_{X|X>1}(x|x>1) dx$$

= $\frac{1}{1-F(1)} \int_{1}^{10} \frac{xf(x)}{1-F(1)} dx$
= $\frac{1}{1-F(1)} \int_{1}^{10} xf(x) dx$
= $\frac{1}{1-F(1)} \int_{1}^{10} x0.7(0.2x-1)^{6} dx.$

There are two subtasks. First,

$$F(1) = \int_0^1 f(x) \, dx = \int_0^1 0.7(0.2x - 1)^6 \, dx = \left[\frac{1}{2}(0.2x - 1)^7\right]_0^1 = -\frac{1}{2}0.8^7 + \frac{1}{2}.$$

Second, using integration by parts,

$$\int x0.7(0.2x-1)^6 dx = \frac{1}{2}x(0.2x-1)^7 - \int \frac{1}{2}(0.2x-1)^7 dx$$
$$= \frac{1}{2}x(0.2x-1)^7 - \frac{1}{3.2}(0.2x-1)^8 + C.$$

Therefore

$$\begin{split} E[X|X>1] &= \frac{1}{1+\frac{1}{2}0.8^7 - \frac{1}{2}} \left[\frac{1}{2}x(0.2x-1)^7 - \frac{1}{3.2}(0.2x-1)^8 \right]_1^{10} \\ &= \frac{1}{1+0.8^7} \left[x(0.2x-1)^7 - \frac{1}{1.6}(0.2x-1)^8 \right]_1^{10} \\ &= \frac{10 - \frac{1}{1.6} + 0.8^7 + \frac{1}{1.6}0.8^8}{1+0.8^7}. \end{split}$$

[By the way, that's approximately 8.01. So surviving the first year greatly increases the expected life span, from 5 to 8 years. Similar remarks apply to the life spans of humans.]